

General Instruction: This is a closed exam, with no any text, notes and electronic devices. The exam lasts 180 minutes. Answer all questions. Print your answer, your name, number the pages and number the problems on provided exam papers. Write on one side only. Please use only black pens or pencils.

Notation of Multiple Linear Regression Model (1): $Y = X\beta + \varepsilon$ where

$$Y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix} \quad X = \begin{pmatrix} 1 & x_{11} & \dots & x_{1k} \\ 1 & x_{21} & \dots & x_{2k} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n1} & \dots & x_{nk} \end{pmatrix} \quad \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_k \end{pmatrix} \quad \varepsilon = \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

$$E(\varepsilon) = \mathbf{0} \quad \varepsilon \sim N(\mathbf{0}, \sigma^2 I) \quad \text{Cov}(\varepsilon) = \sigma^2 I$$

1. For Model (1) (10)
- List all the parameters and their Least Square Estimates (LSE).
 - List all the Maximum Likelihood Estimates (MLE). Are these estimates unbiased?
 - What are the relations of LSE and MLE?
 - Is it necessary to require y_1, y_2, \dots, y_n to be independent?

2. Given the model, $Y = X\beta + \varepsilon$, where $\varepsilon \sim N(\mathbf{0}, \sigma^2 I)$, $Y_{n \times 1}$, $X_{n \times (k+1)}$, $\beta_{(k+1) \times 1}$, $p = k+1$ and $\varepsilon_{n \times 1}$. Prove the following: (25)

a. $b = \hat{\beta} = (XX')^{-1}XY \sim N(\beta, \sigma^2(XX')^{-1})$.

b. $\frac{(\hat{\beta} - \beta)'(XX')^{-1}(\hat{\beta} - \beta)}{\sigma^2} \sim \chi_p^2$.

c. $\hat{\beta}$ and s^2 are independent

d. $(n-p)s^2/\sigma^2 \sim \chi_{n-p}^2$.

e. \bar{Y} and $\sum_{i=1}^n (y_i - \bar{y})^2$ are independent.

3. For model $Y = X\beta + \varepsilon$ with following partitions: (20)

$$Y = \begin{pmatrix} Y^{(1)} & Y^{(2)} \end{pmatrix}, \quad X = \begin{pmatrix} X^{(1)} & X^{(2)} \end{pmatrix}, \quad \beta' = \begin{pmatrix} \beta^{(1)'} & \beta^{(2)'} \end{pmatrix}, \quad \beta^{(1)'} = (\beta_1, \dots, \beta_r), \quad \text{and} \quad \beta^{(2)'} = (\beta_{r+1}, \dots, \beta_p)$$

- a. For test $H_0: \beta^{(2)} = \mathbf{0}$ $H_a: \beta^{(2)} \neq \mathbf{0}$ find the C, γ and m in the following test statistic.

$$\frac{(C\hat{\beta} - \gamma)' [C(XX')^{-1}C']^{-1} (C\hat{\beta} - \gamma)}{Y' [I - X(XX')^{-1}X'] Y} \left(\frac{n-k-1}{m} \right) \sim F_{m, n-p}$$

b. Given that

$$(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_H)'(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}_H) - (\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}})'(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}) = (\mathbf{X}\hat{\boldsymbol{\beta}} - \mathbf{X}\hat{\boldsymbol{\beta}}_H)'(\mathbf{X}\hat{\boldsymbol{\beta}} - \mathbf{X}\hat{\boldsymbol{\beta}}_H)$$

Prove

$$(\mathbf{X}\hat{\boldsymbol{\beta}} - \mathbf{X}\hat{\boldsymbol{\beta}}_H)'(\mathbf{X}\hat{\boldsymbol{\beta}} - \mathbf{X}\hat{\boldsymbol{\beta}}_H) = (\mathbf{C}\hat{\boldsymbol{\beta}} - \boldsymbol{\gamma})' \left[\mathbf{C}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{C}' \right]^{-1} (\mathbf{C}\hat{\boldsymbol{\beta}} - \boldsymbol{\gamma})$$

c. Give the SSE_F and SSE_R in their matrix form, and prove that

$$\frac{(\mathbf{C}\hat{\boldsymbol{\beta}} - \boldsymbol{\gamma})' \left[\mathbf{C}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{C}' \right]^{-1} (\mathbf{C}\hat{\boldsymbol{\beta}} - \boldsymbol{\gamma})}{\mathbf{y}' \left[\mathbf{I} - \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}' \right] \mathbf{y}} \left(\frac{n-k-1}{m} \right) = \frac{SSE_F - SSE_R}{SSE_F} \left(\frac{n-k-1}{m} \right) \sim F_{m, n-p}$$

e. Is the test a Maximum Likelihood Ratio test? What is the significance for it to be a Maximum Likelihood Ratio test? Maximum Likelihood Ratio test

4. For Model (1) (20)

- Give the list square estimates of these all parameters in matrix notation.
- Define estimable of linear function of $\boldsymbol{\beta}$.
- State the Gauss-Markov Theorem and prove the theorem.
- Interpret $\hat{\beta}_3$

5. For Model (1) Let $\mathbf{X}_{(i)}$ be the matrix \mathbf{X} with the i th row removed, and $\mathbf{x}_\ell = (1 \quad x_{\ell 1} \quad \cdots \quad x_{\ell k})'$ $\ell = 1, 2, \dots, n$ so that (10)

$$\sum_{\ell=1}^n \mathbf{x}_\ell \mathbf{x}_\ell' - \mathbf{x}_i \mathbf{x}_i' = \mathbf{X}'\mathbf{X} - \mathbf{x}_i \mathbf{x}_i'$$

Prove

$$\mathbf{b} - \mathbf{b}_{(i)} = \frac{(\mathbf{X}'\mathbf{X})^{-1} \mathbf{x}_i e_i}{1 - h_{ii}}$$

6. For Model (1) (15)

- Define the Studentized Residuals e_i^* $i = 1, 2, \dots, n$.
- Show directly (without using indicator variables interpretation) that that Studentized Residual has Student's t distribution with the degrees of freedom $n-k-2$.